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Statistical Basiscs

Probability Spaces and Measures

Random Variables

Cumulative Distribution Functions

Density Functions

Moments

Test Theory

Parameter Estimation

Test Experiments

updated 12.1.11 1.1/43

Statistical Methods and Learning in Computer Vision

Organization

- Lectures every week, Thursday 11.15am
- Exercises every other week, Tuesday or Wednesday 10.15am
- Oral exam at the end of the lecture
- Topics:
 - necessary basics in measure theory and statistics
 - density estimation and sampling methods
 - subspace methods (PCA, ICA, LDA)
 - learning and classification approaches (SVM, NN)
 - optimization (MRF, PDE)



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Chapter 1 Statistical Basics

Measure and Test Theory

Statistical Methods and Learning in Computer Vision SS 2011

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- 2 Probability Spaces and Measures
- **3** Random Variables
- 4 Cumulative Distribution Functions
- 5 Density Functions
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Randomized Experiment

A randomized experiment is a process with unknown result, which can be arbitrarily often repeated.

Example: tossing two dice.

State Space

The state space Ω is the set of all possible outcomes of a randomized experiment.

Example: $\Omega = \{(i, j) : 1 \le i, j \le 6\}.$

Event

An event is a property which can be observed either to hold or not to hold after the experiment is done. It is a subset of Ω .

Example: $A = \{(i, j) \in \Omega : i + j < 9\}$

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Discrete Probability Space

A triple (Ω, \mathcal{A}, P) is called a discrete probability space if

- the state space Ω is not empty and countable.
- \mathcal{A} is the power set $\mathcal{P}(\Omega)$.
- $P: \mathcal{A} \to \mathbb{R}$ is a mapping with the following properties

•
$$P(A) \geq 0 \ (A \in \mathcal{A})$$

- *P*(Ω) = 1
- for each sequence of pairwise distinct sets from A the σ -additivity holds: $P(\sum_{n \in \mathbb{N}} A_n) = \sum_{n \in \mathbb{N}} P(A_n)$

P is called a probability measure.

Dice Example:

•
$$\Omega = \{(i,j) : 1 \le i, j \le 6\}$$

•
$$\mathcal{A} = \mathcal{P}(\Omega) = \{\{(1,1)\}, \{(1,1)(1,2)\}, \{(1,1)(1,2)(1,3)\}, ...\}$$

•
$$P(A) = \frac{\text{number of elements in A}}{36}$$

Probability Space

σ -Algebra

A σ -algebra \mathcal{A} is a system of subsets of Ω if

- $A \in A \Rightarrow A^c \in A$, where $A^c = \Omega A = \{x \in \Omega : x \notin A\}$ means the complement of A.
- For each sequence (A_n) of sets from $\mathcal{A} \cup_{n \in \mathbb{N}} A_n$ lies in \mathcal{A} .

The elements of \mathcal{A} are called events or measurable sets.

A $\sigma\text{-algebra}$ is closed under finite set operations

- $A, B \in \mathcal{A} \Rightarrow A \cap B \in \mathcal{A}$
- $A, B \in \mathcal{A} \Rightarrow A B \in \mathcal{A}$
- $\emptyset \in \mathcal{A}, \Omega \in \mathcal{A}$

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Borel σ algebra

The Borel σ algebra \mathcal{B}^n consists of all finite unions of I^n , which is the set of right half-open intervals in \mathbb{R}^n . $\mathcal{B}^n = \{\sum_{i=1}^k A_i | k \in \mathbb{N}, A_i \in I^n\}.$

 \mathcal{B}^n is very often used in probability theory, since it contains almost any element in $\mathcal{P}(\mathbb{R}^n)$.

Probability Space

Probability Space

A triple (Ω, \mathcal{A}, P) is called a probability space if

- the state space Ω is not empty.
- \mathcal{A} is a σ -algebra over Ω .
- $P: \mathcal{A} \to \mathbb{R}$ is a mapping with the following properties
 - 1 $P(A) \ge 0 \ (A \in A)$ (non-negative)
 - **2** $P(\Omega) = 1$ (normed)
 - Government for each sequence of pairwise disjoint sets from A the σ-additivity holds: P(∑_{n∈N} A_n) = ∑_{n∈N} P(A_n)

P is a probability measure over the σ algebra A.

In the definition of the discrete probability space the σ algebra was specified as the power set over Ω .

If instead of condition 2, $P(\emptyset) = 0$ holds and \mathcal{A} is a system of subsets over Ω , then P is called a measure, (Ω, \mathcal{A}) is called a measurable space and the triple (Ω, \mathcal{A}, P) a measure space.

Every probability measure is a measure.



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Example of a Probability Measure

Point Probability Measure

Let (Ω, \mathcal{A}) be a state space Ω with σ -algebra \mathcal{A} and $\omega \in \Omega$. Then $\mu_{\omega} : \mathcal{A} \to \{0, 1\}$,

$$\mu_{\omega}(A) = egin{cases} \mathsf{1}, & \omega \in A \ \mathsf{0}, & \omega
otin A \end{cases}$$

defines a probability measure.

Proof:

- μ_ω(A) ≥ 0
- $\mu_{\omega}(\Omega) = 1$
- Let $A_1, ..., A_k, ...$ be pairwise disjoint sets in A, then

$$\mu_{\omega}(\sum_{i=1}^{\infty} A_i) = \begin{cases} 1, & \exists i : \omega \in A_i \\ 0, & \text{otherwise} \end{cases} = \sum_{i=1}^{\infty} \mu_{\omega}(A_i)$$

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Summary

- Probability Space (Ω, A, P), A σ-algebra, P probability measure.
- Probability measure
 - $P(A) \ge 0 \ (A \in A)$
 - $P(\Omega) = 1$
 - for each sequence of pairwise distinct sets from A the σ -additivity holds: $P(\sum_{n \in \mathbb{N}} A_n) = \sum_{n \in \mathbb{N}} P(A_n)$
- For a measure, condition 2 is replaced by P(Ø) = 0 and A does not have to be a σ-algebra.
- Every probability measure is a measure.

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Random Variables

We now aim at transferring a probability measure from one σ -algebra to another.

Dice Example:

We have the state space $\Omega = \{(k, l) : 1 \le k, l \le 6\}$ with event system $\mathcal{P}(\Omega)$ and the uniform probability measure $P : P\{(k, l)\} = \frac{1}{36}$, which make up the probability space $(\Omega, \mathcal{P}(\Omega), P)$.

We are interested in the sum of the dice defined by the mapping $T : \Omega \to \Omega'$ with T(k, l) = k + l and $\Omega' = \{2, ..., 12\}$.

The mapping leads to the probability space $(\Omega', \mathcal{P}(\Omega'), P')$. We are interested in the probability measure P'. For $A' \in \mathcal{P}(\Omega')$, P'(A') is understood as the probability of $P(\{(k, l) : T(k, l) \in A'\}) \in P(\Omega)$.

For example: $P'(\{11, 12\}) = P\{(5, 6), (6, 5), (6, 6)\} = P\{(5, 6)\} + P\{(6, 5)\} + P\{(6, 6)\} = \frac{3}{36}.$

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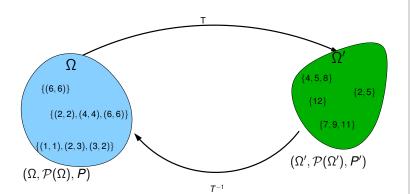
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Preimage Mapping

Let $T : \Omega \to \Omega'$ be an arbitrary mapping. Then the preimage mapping $T^{-1} : \mathcal{P}(\Omega') \to \mathcal{P}(\Omega)$ is defined by

$$T^{-1}(\mathbf{A}') = \{ \omega \in \Omega : T(\omega) \in \mathbf{A}' \}, \mathbf{A}' \in \mathcal{P}(\Omega')$$

The preimage of a σ -algebra is a σ -algebra.



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Random Variable

Let (Ω, \mathcal{A}) and (Ω', \mathcal{A}') be measurable spaces. A mapping $X : (\Omega, \mathcal{A}) \to (\Omega', \mathcal{A}')$ is called $\mathcal{A} - \mathcal{A}'$ measurable if

 $X^{-1}(A') = \{\omega \in \Omega : X(\omega) \in A'\} \in A$

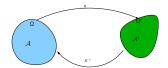
Such a measurable mapping X is called a random variable.

Image Measure

Let $X : (\Omega, \mathcal{A}) \to (\Omega', \mathcal{A}')$ be a random variable and P a measure over \mathcal{A} . Then

$$P'(A') := P_X(A') := P(X^{-1}(A')), A' \in A$$

defines a measure over \mathcal{A}' . Is *P* a probability measure then P_X is a probability measure over \mathcal{A}' . P_X is called the image measure of *P* by X.



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Example

Dice Example:

State spaces $\Omega = \{(k, l) : 1 \le k, l \le 6\}$ and $\Omega' = \{2, ..., 12\}$ with power sets as event system and the uniform probability measure on $(\Omega, \mathcal{P}(\Omega))$.

Define the mapping $X : (\Omega, \mathcal{P}(\Omega)) \to (\Omega', \mathcal{P}(\Omega'))$, X((k, I)) = k + I. Is X a random variable? It holds that $X^{-1}(A') = \{\omega \in \Omega : X(\omega) \in A'\} \in \mathcal{P}(\Omega)$ since for every value in Ω' we can find two dice results which sum up to this value. It follows that X is $\mathcal{P}(\Omega) - \mathcal{P}(\Omega')$ -measurable and, thus, a random variable.

Furthermore, *P* is a probability measure. Hence, the image measure P_X under *X* which is defined by $P_X(A') = P(X^{-1}(A'))$ is a probability measure on $(\Omega', \mathcal{P}(\Omega'))$.

For example

$$P_X\{2,4,5\} = P(X^{-1}\{2,4,5\}) = P\{(k,l) : X(k,l) \in \{2,4,5\}\} = P\{(1,1), (2,2), (1,3), (3,1), (2,3), (3,2)\} = \frac{6}{36} = \frac{1}{6}.$$

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Probability Distribution

The notion 'random variable' is just a name. It is neither a variable nor is it random, but a measurable mapping. By means of a random variable measures can be transferred from one σ -algebra to another.

Probability Distribution

Let $X : (\Omega, \mathcal{A}, P) \to (\Omega', \mathcal{A}')$ be a random variable. Then the image measure P_X of P by X is called probability distribution.

Every probability measure can be understood as a distribution, since there is always a random variable having *P* as its image measure: the identical mapping $Id : \Omega \rightarrow \Omega$. Hence, the notions probability distribution and probability measure are often used equivalently.

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 A random variable is a measurable mapping, which transfers probability measures from one space to another

- The image measure of a random variable is called probability distribution
- Every probability measure is a probability distribution by means of the random variable defined by the identity mapping.

Cumulative Distribution Function

Let *P* be a probability measure on the Borel σ -algebra \mathcal{B} . Then $F_P : \mathbb{R} \to \mathbb{R}$ is called cumulative distribution function of *P* if

$$F_P(x) = P((-\infty, x)), x \in \mathbb{R}$$

E.

Cumulative Distribution Ect.

 $F_{\omega}(x) = \begin{cases} 0, & x \leq \omega \\ 1, & x > \omega \end{cases}$

Example:

Point Measure on B

 $\mu_{\omega}(\mathcal{A}) = egin{cases} \mathsf{0}, & \omega
otin \mathcal{A} \ \mathsf{1}, & \omega \in \mathcal{A} \end{cases}$

Note: a probability distribution is a probability measure defined on a specific system of sets (a σ -algebra). The cumulative distribution function is defined for points in \mathbb{R} .

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Cumulative Distribution Function

$$F_P(x) = P((-\infty, x)), x \in \mathbb{R}$$

Properties of Cumulative Distribution Functions

The cumulative distribution function F_P of a probability measure P over \mathcal{B} has the following properties

- *F_P* is monotonously increasing
- *F_P* is left-continuous
- $\lim_{x\to -\infty} F_P(x) = 0$
- $\lim_{x\to\infty} F_P(x) = 1$

Each function with these properties uniquely describes a probability measure.

Each probability measure is defined uniquely by its distribution function.

$$P([a,b)) = P(-\infty,b) - P(-\infty,a) = F_P(b) - F_P(a)$$

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- The cumulative distribution function is defined as: $F_P(x) = P((-\infty, x)), x \in \mathbb{R}$
- The probability of an element of the Borel *σ*-algebra *B* can be determined by P([a, b)) = F_P(b) - F_P(a)
- A probability measure and its cumulative distribution function uniquely determine each other.

Density Function

Let $F_P : \mathbb{R} \to \mathbb{R}$ be the cumulative distribution function of a probability measure *P* over *B*. A measurable function $f : \mathbb{R} \to \mathbb{R}_+ \cup \{\infty\}$ is called a density function if and only if

$$F_P(t) = \int_{-\infty}^t f(x) \, \mathrm{d}x, \ t \in \mathbb{R}$$

Densities are not unique. However, different densities of the same probability measure differ only over null sets.

Relationship between probability measures, cumulative distribution functions and density functions:

$$P([a,b)) = P((-\infty,b)) - P((-\infty,a))$$

= $F_P(b) - F_P(a)$
= $\int_{-\infty}^{b} f_P(x) dx - \int_{-\infty}^{a} f_P(x) dx$
= $\int_{a}^{b} f_P(x) dx$

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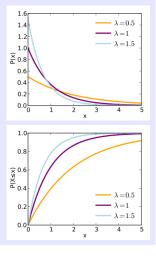
Exponential Distribution

Density Function

$$f_{\lambda}(x) = egin{cases} \lambda \exp^{-\lambda x} & x \geq 0 \ 0 & x < 0 \end{cases}$$

Cumulative Distribution Function

$$F_{\lambda}(x) = \int_{-\infty}^{x} f_{\lambda}(t) dt = 1 - \exp^{-\lambda x}$$



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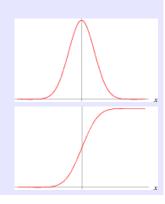
Normal Distribution

Density Function

$$f(x)=\frac{1}{\sqrt{2\pi}}\exp^{-\frac{x^2}{2}}, x\in\mathbb{R}$$

Cumulative Distribution Function

$$F(t)=rac{1}{\sqrt{2\pi}}\int_{-\infty}^t \exp^{-rac{x^2}{2}}\,\mathrm{d}x,\ t\in\mathbb{R}$$



Density of the Image Measure

We now want to transfer the density of a probability measure P from one σ -algebra to another.

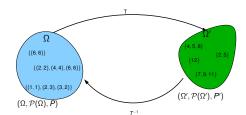
Let *P* be a probability measure over \mathcal{B}^n with density *f* and $T : (\mathbb{R}^n, \mathcal{B}^n) \to (\mathbb{R}^n, \mathcal{B}^n)$ a measurable mapping. When does a density of the image measure P_T exist?

We have

$$P_T(B') = P(T^{-1}(B')) = \int_{T^{-1}(B')} f \, dx \, (B' \in \mathcal{B}^n)$$

Does a density $g : \mathbb{R}^n \to \mathbb{R}_+$ of the image measure P_T exist with

$$P(T^{-1}(B')) = \int_{T^{-1}(B')} f \, \mathrm{d}x = \int_{B'} g \, \mathrm{d}x = P_T(B'), \text{ for } B' \in \mathcal{B}^n$$
?





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Density of the Image Measure: Density Transformation Theorem

Let *P* be a probability measure over \mathcal{B}^n with a density function *f*, and $T : \mathbb{R}^n \to \mathbb{R}^n$ a measurable mapping for which holds

- T is continuously differentiable.
- The functional determinant $\Delta_T(x) = \frac{\partial T}{\partial x} \neq 0$ for $x \in \mathbb{R}^n$.
- *T* is injective meaning that the inverse mapping $T^{-1}: T(\mathbb{R}^n) \to \mathbb{R}^n$ exists.

Then T^{-1} is measurable and a density function g of the image measure P_T is given by

 $g(y) = egin{cases} rac{f(T^{-1}(y))}{|\Delta_T(T^{-1}(y))|} & y \in \mathcal{T}(\mathbb{R}^n) \ 0 & y \in \mathbb{R}^n - \mathcal{T}(\mathbb{R}^n) \end{cases}$



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Density Transformation Example

Let *X* be a random variable distributed according to the standard normal distribution, i.e. $P_X = \mathcal{N}(0, 1)$, meaning that the density function of *X* is given by

$$f(y) = \frac{1}{\sqrt{2\pi}} \exp\{-\frac{y^2}{2}\}$$

Let $\mu, \sigma \in \mathbb{R}, \sigma > 0$. The mapping

$$T(\mathbf{x}) = \sigma \mathbf{x} + \mu$$

is measurable. It is continuously differentiable, $\Delta_T(x) = \sigma > 0$ and the inverse transform exists $T^{-1}(y) = \frac{y-\mu}{\sigma}$. Hence, a density of the image measure P_T is given by

$$g(y) = \frac{f(T^{-1}(y))}{|\Delta_T(T^{-1}(y))|} = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(y-\mu)^2}{2\sigma^2}\right\}$$

g is the density of a probability measure, namely of the normal distribution $\mathcal{N}(\mu, \sigma^2)$.

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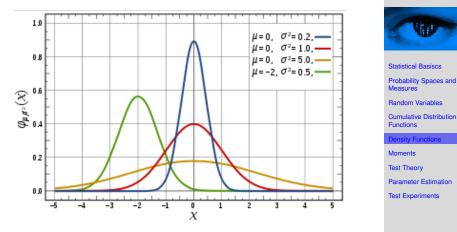
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The Normal Distribution



The smaller σ^2 the steeper rises the density function. It is centered at μ .

90% of the area under curve lie within the interval $\mu \pm 1.6\sigma$. 95% of the area under curve lie within the interval $\mu \pm 2\sigma$. 99% of the area under curve lie within the interval $\mu \pm 3\sigma$.

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- The density *f* of a probability measure *P* is defined by $F_P(t) = \int_{-\infty}^t f(x) dx, \ t \in \mathbb{R}$
- $P([a,b)) = \int_a^b f(x) \,\mathrm{d}x$
- A density of an image measure *P_X* under a random variable *X* is given by the transformation theorem.
- The density of the normal distribution $\mathcal{N}(\mu, \sigma)$ is given as the transformed standard normal density under the random variable $T(x) = \sigma x + a$

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Expectation Value

Let (Ω, \mathcal{A}, P) be a probability space, *f* a density of P_X and $X : (\Omega, \mathcal{A}) \to (\mathbb{R}, \mathcal{B})$ a random variable. Then the expectation value of the random variable *X* is given by

$$E(X) = \int_{-\infty}^{\infty} x f(x) \, \mathrm{d}x$$

Examples

Exponential Distribution:
$$E(X) = \int_{-\infty}^{\infty} x \ \lambda \exp^{-\lambda x} dx = \frac{1}{\lambda}$$

Normal Distribution: $E(X) = \int_{-\infty}^{\infty} x \ \frac{1}{\sqrt{2\pi\sigma}} \exp\{-\frac{(x-\mu)^2}{2\sigma^2}\} dx = \mu$

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Let (Ω, \mathcal{A}, P) be a probability space, *f* a density of *P* and $X : (\Omega, \mathcal{A}) \to (\mathbb{R}, \mathcal{B})$ a random variable. Then the *k*-th moment of the random variable *X* is given by

$$E(X^k) := \int_{-\infty}^{\infty} x^k f(x) \, \mathrm{d}x.$$

The k-th centralized moment is given by

$$E[(X - E(X))^k] := \int_{-\infty}^{\infty} (x - E(x))^k f(x) dx$$

The first moment is the expectation value, the 0-th equals 1. The second centralized moment is called the variance $V_P(X)$. $\sqrt{V_P(X)}$ is called the standard deviation. The *k*-th (centralized) moment of the probability measure *P* is understood as that of the identity random variable Id_B .

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Properties of the Expectation Value

Let $\lambda \in \mathbb{R}$, $A \in \mathbb{R}^{n \times n}$, then

$$E(c) = c$$

$$E(\lambda X) = \lambda E(X)$$

$$E(X + Y) = E(X) + E(Y)$$

$$E(AX) = AE(X)$$

Translation Theorem for the Variance

Let (Ω, \mathcal{A}, P) be a probability space and X a random variable: $V(X) = E(X^2) - [E(X)]^2$

Proof

$$V(X) = E[(X - E(X))^{2}]$$

= $E[X^{2} - 2XE(X) + E(X)^{2}]$
= $E(X^{2}) - 2E(X)E(X) + E(X)^{2}$
= $E(X^{2}) - E(X)^{2}$

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Test Theory Parameter Estimation Test Experiments The variance measures the expected squared deviation of a random variable from its expectation value.

Covariance and Correlation Coefficient

Let $X, Y : (\Omega, \mathcal{A}, P) \to (\mathbb{R}, \mathcal{B})$ be two random variables with variance V(X), V(Y). Then

• The covariance of X and Y is defined as

Cov(X, Y) := E[(X - E(X))(Y - E(Y))]

• The correlation coefficient of X and Y is defined as

 $Cor(X, Y) := \frac{Cov(X, Y)}{\sqrt{V(X)}\sqrt{V(Y)}}$

If the covariance is 0, X and Y are called uncorrelated. The larger the covariance the more vary X and Y together. The correlation coefficient is a normalized version of the covariance and lies between -1 and 1.



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Moment

Test Theory

Parameter Estimation

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Corollary

- Cov(X, Y) = E(XY) E(X)E(Y)
- Cov(aX + c, bY + d) = abCov(X, Y)
- Cov(X, X) = V(X)
- $V(aX+c) = a^2 V(X)$
- V(X + Y) = V(X) + V(Y) + 2Cov(X, Y)

Uncorrelated Random Variables

For two random variables X and Y the following statements are equivalent

- Cor(X, Y) = 0 (X and Y uncorrelated)
- Cov(X, Y) = 0
- E(XY) = E(X)E(Y)
- V(X + Y) = V(X) + V(Y) (equation by Bienaymé)

Summary

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•
$$E_P(X) = \int_{-\infty}^{\infty} x f(x) dx$$

• $V(X) = E(X - E(X))^2 = E(X^2) - E(X)^2$

•
$$Cov(X, Y) := E[(X - E(X))(Y - E(Y))]$$

•
$$Cor(X, Y) := \frac{Cov(X, Y)}{\sqrt{V(X)}\sqrt{V(Y)}}$$

• $Cov(X, Y) = 0 \Leftrightarrow E(XY) = E(X)E(Y) \Leftrightarrow V(X + Y) = V(X) + V(Y)$

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Independence

Let (Ω, \mathcal{A}, P) be a probability space and $(X_i)_{i \in \mathbb{N}}$ a set of random variables $X_i : (\Omega, \mathcal{A}) \to (\Omega_i, \mathcal{A}_i)$. The random variables X_i are called stochastically independent, if and only if

$$P(\{\omega \in \Omega | X_1(\omega) \in A_1, ..., X_n(\omega) \in A_n\}) = \prod_{i=1}^n P(\{\omega \in \Omega | X_i(\omega) \in A_i\}), A_i \in A_i$$

Independent random variables are uncorrelated, but uncorrelated random variables are not necessarily independent.



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Marginal Density

Let X = (Y, Z) be a random variable and f a density of P_x on \mathbb{R}^2 . Then both P_Y and P_Z have marginal densities on $(\mathbb{R}, \mathcal{B})$:

$$f_Y(y) = \int_{-\infty}^{\infty} f(y, z) dz$$
 $f_Z(z) = \int_{-\infty}^{\infty} f(y, z) dy$

Y and *Z* are independent if and only if $f(y, z) = f_Y(y)f_Z(z)$. In general the joint density *f* cannot be recovered from the marginal distributions only.

Conditional Density

Let X = (Y, Z) be a random variable with density and f a density of P_X on \mathbb{R}^2 with marginal density $f_Y(y) \neq 0$. Then the density

$$f_c(z|Y=y) = \frac{f(y,z)}{f_Y(y)}$$

is called the conditional density of Z given Y = y.

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Statistical Experiment

Let $(\mathbb{H}, \mathcal{H})$ be a measurable space, the sample space, and \mathcal{W} a set of possible probability measures. The set $(\mathbb{H}, \mathcal{H}, \mathcal{W})$ is called statistical experiment if there is exactly one among the probability measures in \mathcal{W} which manifests itself through its sample realizations $x \in \mathbb{H}$.

Statistic

Let $(\mathbb{H}, \mathcal{H})$ be a sample space and $(\mathbb{D}, \mathcal{D})$ a measurable space. An \mathcal{H} - \mathcal{D} -measurable mapping $\mathcal{T} : \mathbb{H} \to \mathcal{D}$ is called statistic.

Test Problem

Let $(\mathbb{H}, \mathcal{H}, \mathcal{W})$ be a statistical experiment. In test theory, \mathcal{W} is partitioned into two subsets: $\mathcal{W}_1 = \{\overline{P}\}$ is called the hypothesis, $\mathcal{W}_2 = \mathcal{W} - \mathcal{W}_1$ the alternative. For a given sample $x \in \mathbb{H}$ we examine if x is distributed according to the distribution in \mathcal{W}_1 or not. A test is defined as a function $\varphi : \mathbb{H} \to \{0, 1\}$ for which

 $\varphi(x) = 1$ if the hypothesis is rejected and $\varphi(x) = 0$ otherwise.

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Maximum Likelihood Estimation

Let $(\mathbb{H}, \mathcal{H}, \mathcal{W})$ be a statistical experiment, where \mathcal{W} denotes a family of probability measures parameterized by Γ , $\mathcal{W} = \{P_{\gamma} | \gamma \in \Gamma\}$. We want to find out the parameter describing the true probability measure behind a given sample realization. An estimator T is a mapping from the sample space \mathbb{H} to the set of parameters Γ , which assigns to each sample realization $x \in \mathbb{H}$ an estimated value of the unknown parameter γ .

Likelihood Function

For $x \in \mathbb{H}$ the mapping $L_x : \gamma \to \mathbb{R}^+$, $L_x(\gamma) = \prod_{i=1}^n f_{\gamma}(x_i), \gamma \in \Gamma$ is called the likelihood function for the sample realization x.

Maximum Likelihood Estimator

A maximum likelihood estimator T is an estimator, which for a given sample realization $x \in \mathbb{H}$ finds the most likely probability distribution P_{γ} by maximizing the likelihood function L_x :

Find
$$T(x)$$
 s.t. $L_x(T(x)) = \sup_{\gamma \in \Gamma} L_x(\gamma)$

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Maximum A Posteriori

If additionally a prior over the parameters Γ is given, then we obtain a maximum a posteriori estimator.

Maximum A Posteriori

Maximum Likelihood:

$$\hat{\gamma}_{ML}(x) = \operatorname*{argmax}_{\gamma \in \Gamma} \prod_{i=1}^{n} f_{\gamma}(x_i)$$

Maximum A Posteriori:

$$\hat{\gamma}_{MAP}(x) = \operatorname*{argmax}_{\gamma \in \Gamma} \prod_{i=1}^{n} \frac{f_{\gamma}(x_{i})p(\gamma)}{\int_{\gamma \in \Gamma} f_{\gamma}(x_{i})p(\gamma)d\gamma}$$

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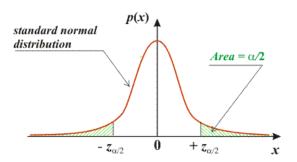
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Test Experiment

Let $(\mathbb{H},\mathcal{H},\mathcal{W}=\mathcal{W}_1\cup\mathcal{W}_2)$ be a statistical experiment. A statistical test proceeds as follows

- Define a test statistic T with known image measure P
 T
 under the hypothesis.
- Compute the critical region for values of *T* under the hypothesis using the distribution specific quantiles.
- Decide to either fail to reject the hypothesis or reject the hypothesis.



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 χ^2 -Test

The χ^2 test examines if a given sample $X = (X_1, ..., X_n)$ is distributed according to a probability measure *P*.

χ^2 -Test

Let $(A_1, ..., A_k)$ be a partition of \mathbb{H} into *k* disjoint subsets. Then we compute the random vector

$$Y_i = \sum_{j=1}^n \mathbf{1}_{A_i} \circ X_j,$$

which counts the number of samples falling into A_i . The expected value of Y_i under the hypothesis distribution P is

$$E_P(Y_i) = \sum_{j=1}^n E_P(1_{A_i} \circ X_j) = nP_X(A_i)$$

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χ^2 -Test

The Pearson Test Statistic measures the deviation of the expected value of Y_i from the observed value. Small values support the hypothesis.

$$T(x) = \sum_{i=1}^{n} \frac{(Y_i(x) - nP_X(A_i))^2}{nP_X(A_i)}$$

The statistic *T* is asymptotically distributed according to the $\chi^2(k-1)$ distribution. Let *t* be the $(1 - \alpha)$ -quantile of this distribution. Then

$$P(\{x \in \mathbb{H} : T(x) \le t\}) = 1 - \alpha$$

If T(x) > t we reject the hypothesis, otherwise we do not reject the hypothesis.



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P-Values

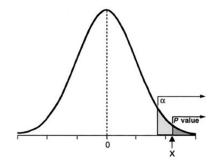
Sometimes we do not want a binary decision against a hypothesis.

P-Values

A p-value is the minimum significance level α for which the null-hypothesis is rejected given the value of the test statistic T(x).

$$v(X) = \inf\{\alpha \in [0, 1] : \varphi_{\alpha}(T(X)) = 1\}$$

The larger v(X) the more conform is the value T(X) with the assumed probability measure.



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