# 8. Clustering

#### **Motivation**

- Supervised learning is good for interaction with humans, but labels from a supervisor are sometimes hard to obtain
- Clustering is unsupervised learning, i.e. it tries to learn only from the data
- Main idea: find a similarity measure and group similar data objects together
- Clustering is a very old research field, many approaches have been suggested
- Main problem in most methods: how to find a good number of clusters





#### **Categories of Learning**

Learning

#### Unsupervised Learning

clustering, density estimation

Supervised Learning

learning from a training data set, inference on the test data

Reinforcement Learning

no supervision, but a reward function

In unsupervised learning, there is no ground truth information given.

Most Unsupervised Learning methods are based on **Clustering**.



- Given: data set  $\{x_1, \dots, x_N\}$ , number of clusters K
- Goal: find cluster centers  $\{\mu_1,\ldots,\mu_K\}$  so that

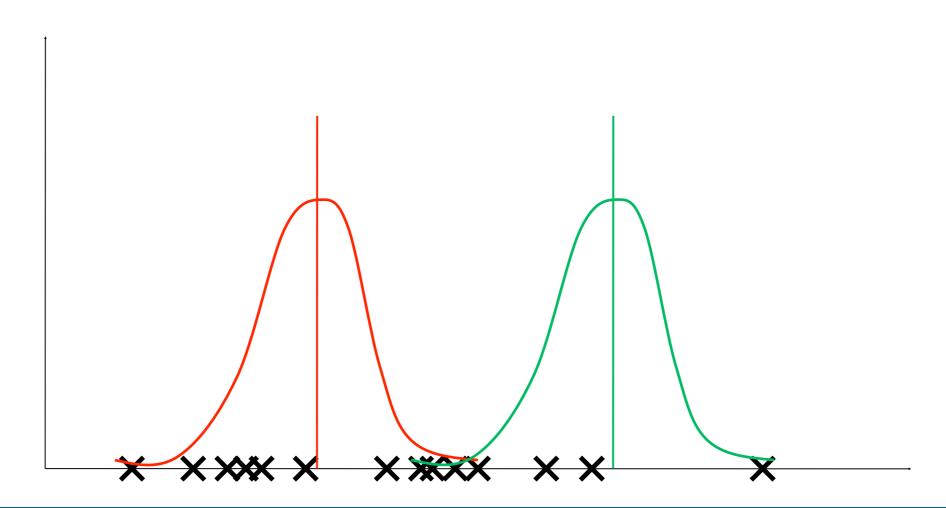
$$J = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2$$

is minimal, where  $r_{nk}=1$  if  $\mathbf{x}_n$  is assigned to  $\boldsymbol{\mu}_k$ 

- Idea: compute  $r_{nk}$  and  $\mu_k$  iteratively
- Start with some values for the cluster centers
- Find optimal assignments  $r_{nk}$
- Update cluster centers using these assignments
- Repeat until assignments or centers don't change

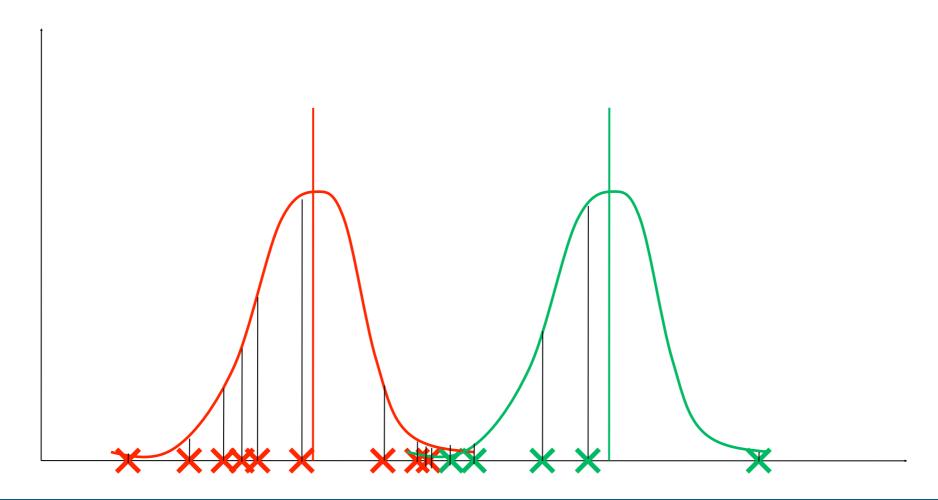


Initialize cluster means:  $\{oldsymbol{\mu}_1,\dots,oldsymbol{\mu}_K\}$ 



Find optimal assignments:

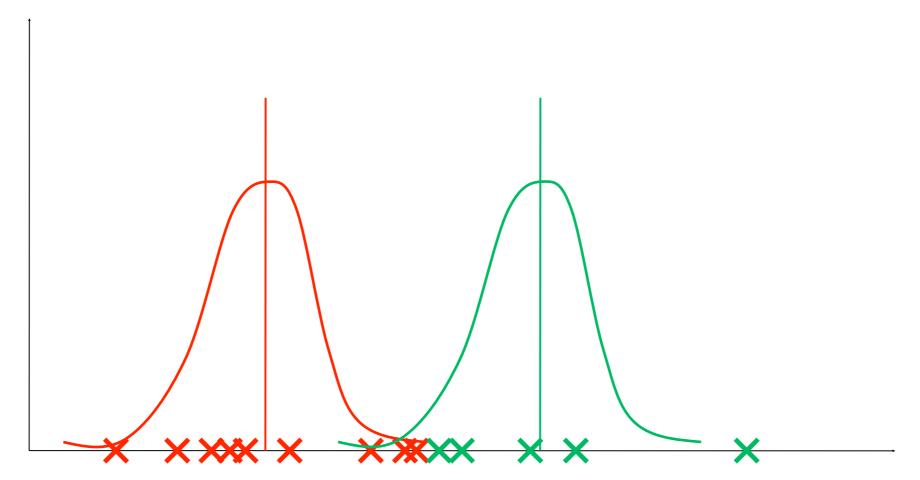
$$r_{nk} = \begin{cases} 1 & \text{if } k = \arg\min_{j} \|\mathbf{x}_n - \boldsymbol{\mu}_j\| \\ 0 & \text{otherwise} \end{cases}$$



Find new optimal means:

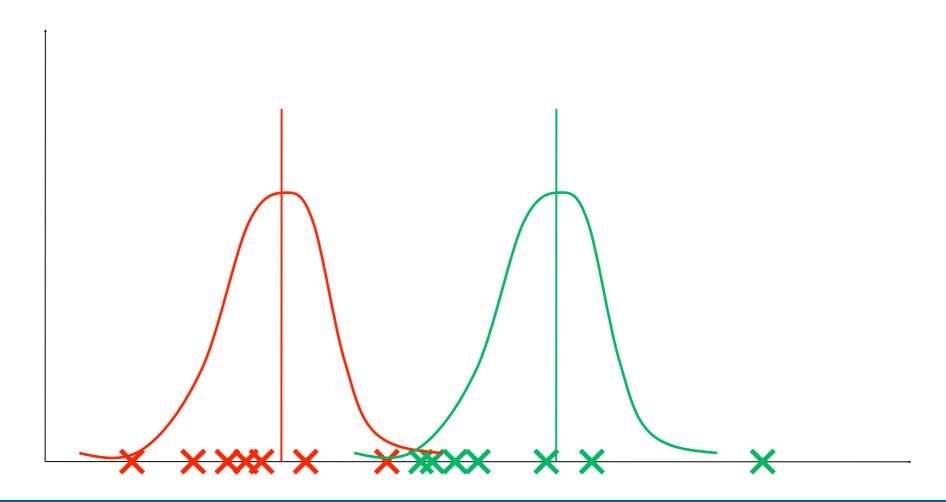
$$\frac{\partial J}{\partial \boldsymbol{\mu}_k} = 2\sum_{n=1}^N r_{nk}(\mathbf{x}_n - \boldsymbol{\mu}_k) \stackrel{!}{=} 0$$

$$\Rightarrow \boldsymbol{\mu}_k = \frac{\sum_{n=1}^N r_{nk} \mathbf{x}_n}{\sum_{n=1}^N r_{nk}}$$

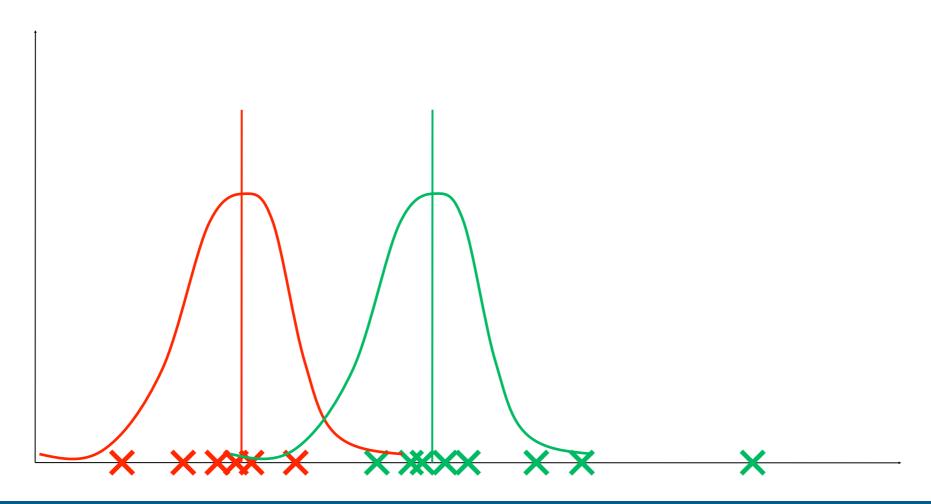


Find new optimal assignments:

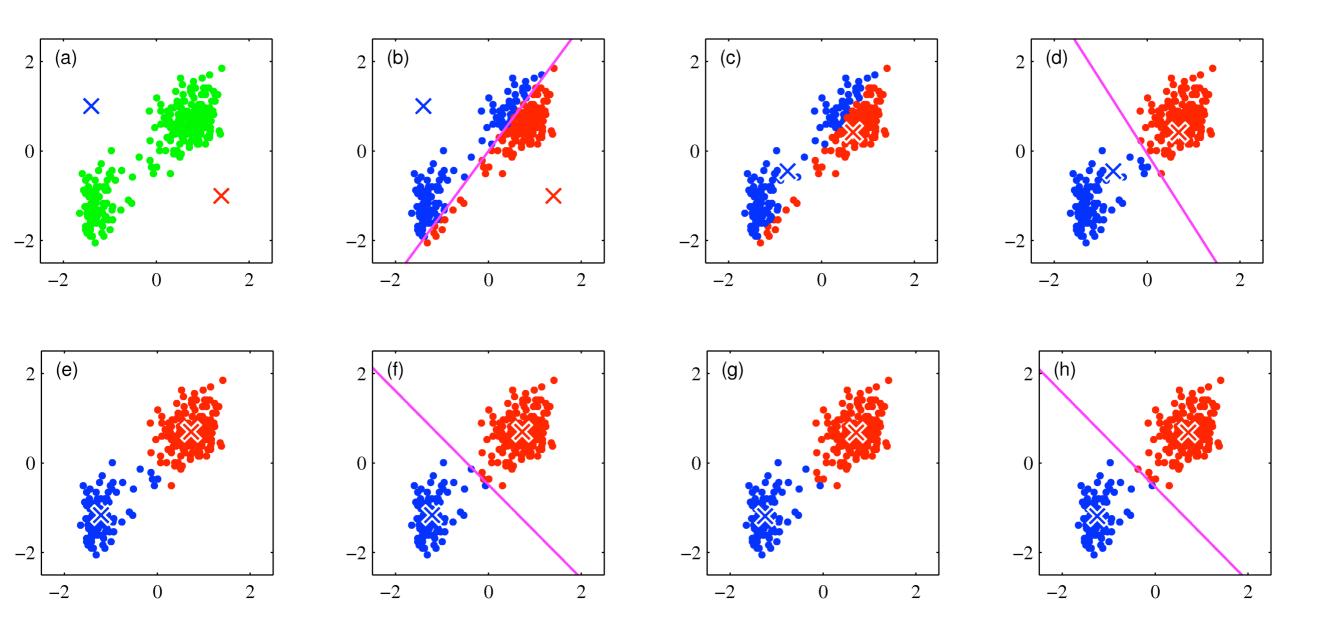
$$r_{nk} = \begin{cases} 1 & \text{if } k = \arg\min_{j} \|\mathbf{x}_n - \boldsymbol{\mu}_j\| \\ 0 & \text{otherwise} \end{cases}$$



Iterate these steps until means and assignments do not change any more



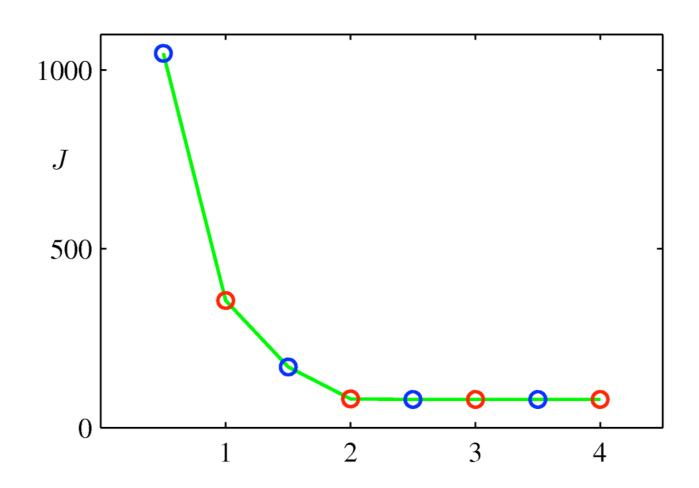
#### 2D Example



- Real data set
- Random initialization

 Magenta line is "decision boundary"

#### **The Cost Function**



- After every step the cost function J is minimized
- Blue steps: update assignments
- Red steps: update means
- Convergence after 4 rounds





## K-means for Segmentation





K = 3



K = 10



Original image











#### K-Means: Additional Remarks

- K-means converges always, but the minimum is not guaranteed to be a global one
- There is an **online** version of K-means
  - After each addition of  $\mathbf{x}_n$ , the nearest center  $\boldsymbol{\mu}_k$  is updated:  $\boldsymbol{\mu}_k^{\mathrm{new}} = \boldsymbol{\mu}_k^{\mathrm{old}} + \eta_n(\mathbf{x}_n \boldsymbol{\mu}_k^{\mathrm{old}})$
- The **K-medoid** variant:
  - Replace the Euclidean distance by a general measure
     V.

$$\tilde{J} = \sum_{n=1}^{N} \sum_{k=1}^{K} r_{nk} \mathcal{V}(\mathbf{x}_n, \boldsymbol{\mu}_k)$$



#### **Mixtures of Gaussians**

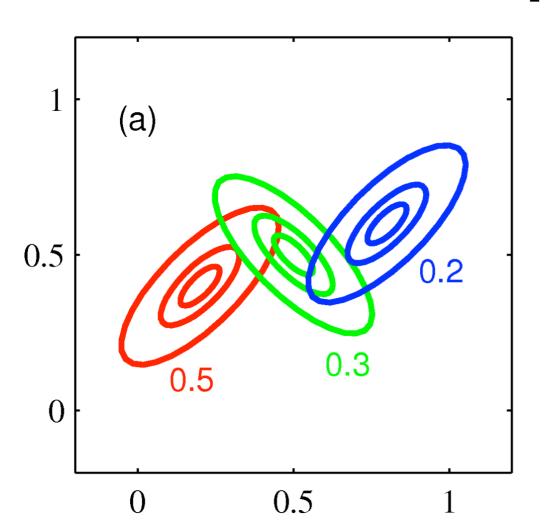
- Assume that the data consists of K clusters
- The data within each cluster is Gaussian
- For any data point  $\mathbf{x}$  we introduce a K-dimensional binary random variable  $\mathbf{z}$  so that:

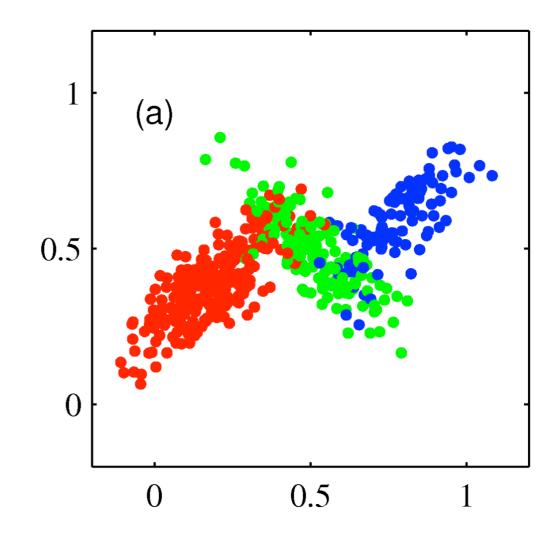
$$p(\mathbf{x}) = \sum_{k=1}^{K} \underbrace{p(z_k = 1)}_{=:\pi_k} \mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

where

$$z_k \in \{0, 1\}, \quad \sum_{k=1}^K z_k = 1$$

#### A Simple Example





- Mixture of three Gaussians with mixing coefficients
- Left: all three Gaussians as contour plot
- Right: samples from the mixture model, the red component has the most samples



#### **Parameter Estimation**

• From a given set of training data  $\{\mathbf{x}_1, \dots, \mathbf{x}_N\}$  we want to find parameters  $(\pi_{1,\dots,K}, \boldsymbol{\mu}_{1,\dots,K}, \Sigma_{1,\dots,K})$  so that the likelihood is maximized (MLE):

$$p(\mathbf{x}_1, \dots, \mathbf{x}_N \mid \pi_{1,\dots,K}, \boldsymbol{\mu}_{1,\dots,K}, \Sigma_{1,\dots,K}) = \prod_{n=1}^N \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \Sigma_k)$$

or, applying the logarithm:

$$\log p(X \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^{N} \log \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

 However: this is not as easy as maximumlikelihood for single Gaussians!



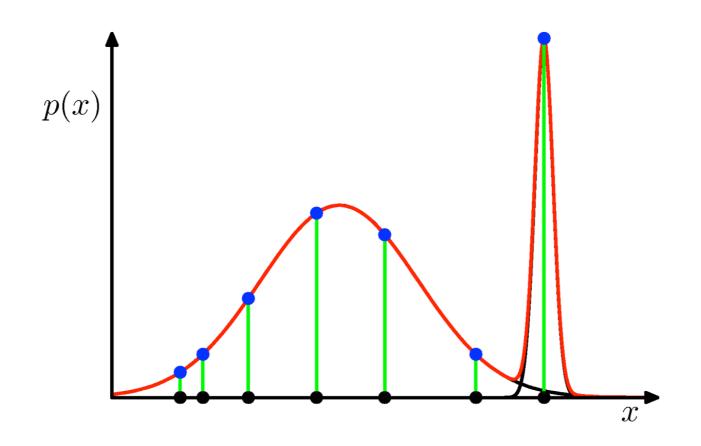
#### **Problems with MLE for Gaussian Mixtures**

- Assume that for one k the mean  $\mu_k$  is exactly at a data point  $\mathbf{x}_n$ 
  - For simplicity: assume that  $\Sigma_k = \sigma_k^2 I$
  - Then:  $\mathcal{N}(\mathbf{x}_n \mid \mathbf{x}_n, \sigma_k^2 I) = \frac{1}{\sqrt{2\pi}\sigma_k^D}$
  - This means that the overall log-likelihood can be maximized arbitrarily by letting  $\sigma_k \to 0$  (overfitting)
- Another problem is the identifiability:
  - The order of the Gaussians is not fixed, therefore:
  - There are K! equivalent solutions to the MLE problem





#### Overfitting with MLE for Gaussian Mixtures



- One Gaussian fits exactly to one data point
- It has a very small variance, i.e. contributes strongly to the overall likelihood
- In standard MLE, there is no way to avoid this!



#### **Expectation-Maximization**

- EM is an elegant and powerful method for MLE problems with latent variables
- Main idea: model parameters and latent variables are estimated iteratively, where average over the latent variables (expectation)
- A typical example application of EM is the Gaussian Mixture model (GMM)
- However, EM has many other applications
- First, we consider EM for GMMs



• First, we define the responsibilities:

$$\gamma(z_{nk}) = p(z_{nk} = 1 \mid \mathbf{x}_n) \qquad z_{nk} \in \{0, 1\}$$
$$\sum_{k} z_{nk} = 1$$

• First, we define the responsibilities:

$$\gamma(z_{nk}) = p(z_{nk} = 1 \mid \mathbf{x}_n)$$

$$= \frac{\pi_k \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{i=1}^K \pi_i \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_i, \boldsymbol{\Sigma}_i)}$$



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ullet Next, we derive the log-likelihood wrt. to  $\mu_k$ :

$$\frac{\partial \log p(X \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})}{\partial \boldsymbol{\mu}_k} \stackrel{!}{=} \mathbf{0}$$

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and we obtain: 
$$\mu_k = \frac{\sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n}{\sum_{n=1}^N \gamma(z_{nk})}$$

We can do the same for the covariances:

$$\frac{\partial \log p(X \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})}{\partial \Sigma_k} \stackrel{!}{=} \mathbf{0}$$

and we obtain:

$$\Sigma_k = \frac{\sum_{n=1}^N \gamma(z_{nk})(\mathbf{x}_n - \boldsymbol{\mu}_k)(\mathbf{x}_n - \boldsymbol{\mu}_k)^T}{\sum_{n=1}^N \gamma(z_{nk})}$$

• Finally, we derive wrt. the mixing coefficients  $\pi_k$ :

$$\frac{\partial \log p(X \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \Sigma)}{\partial \pi_k} \stackrel{!}{=} \mathbf{0}$$
 where:  $\sum_{k=1}^K \pi_k = 1$ 

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and the result is:  $\pi_k = \frac{1}{N} \sum_{i=1}^{N} \gamma(z_{nk})$ 



### **Algorithm Summary**

- 1.Initialize means  $\mu_k$  covariance matrices  $\Sigma_k$  and mixing coefficients  $\pi_k$
- 2.Compute the initial log-likelihood  $\log p(X \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma})$
- 3. E-Step. Compute the responsibilities:

$$\gamma(z_{nk}) = \frac{\pi_k \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)}{\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)}$$

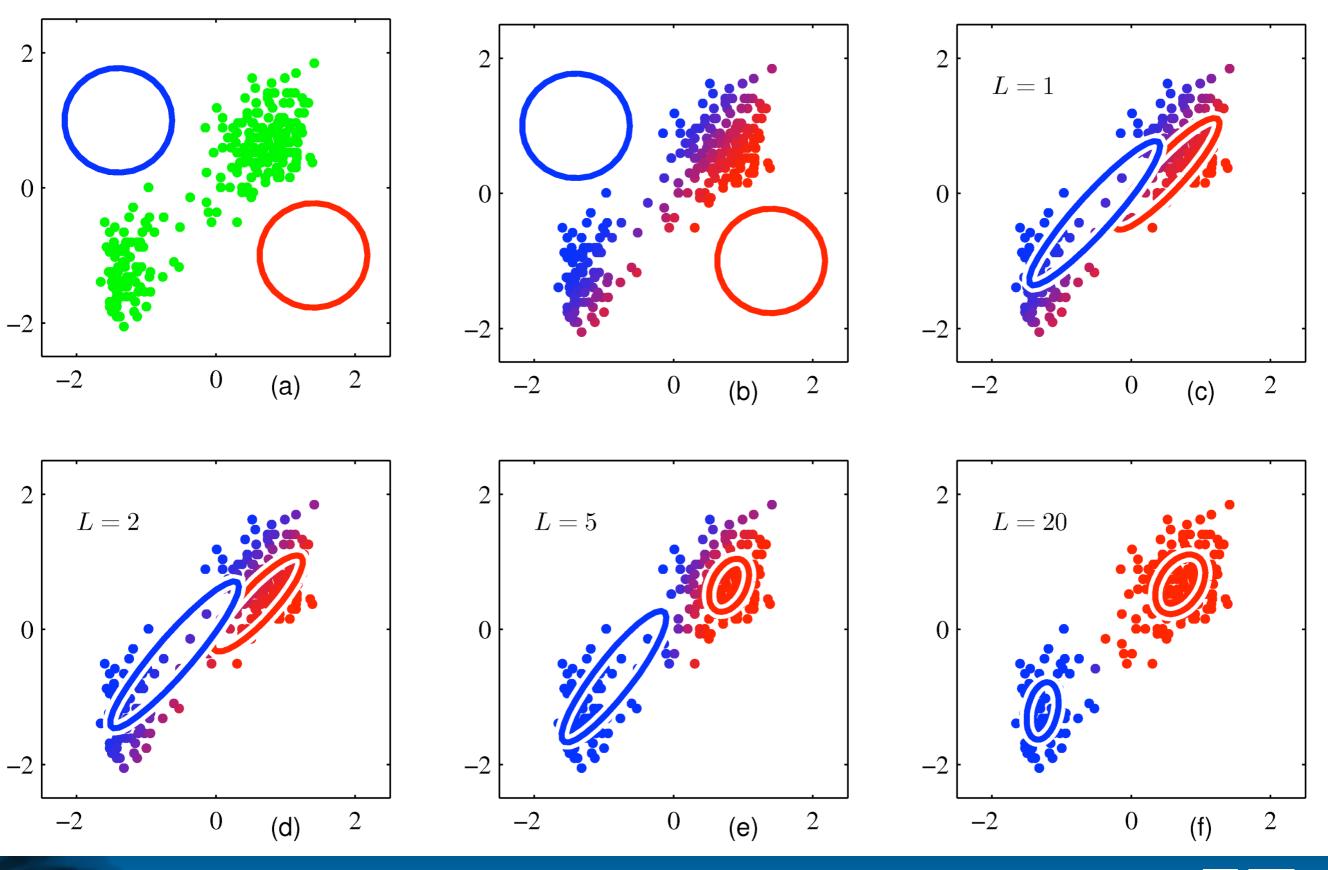
4. M-Step. Update the parameters:

$$\mu_k^{\text{new}} = \frac{\sum_{n=1}^{N} \gamma(z_{nk}) \mathbf{x}_n}{\sum_{n=1}^{N} \gamma(z_{nk})} \quad \Sigma_k^{\text{new}} = \frac{\sum_{n=1}^{N} \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}}) (\mathbf{x}_n - \boldsymbol{\mu}_k^{\text{new}})^T}{\sum_{n=1}^{N} \gamma(z_{nk})} \quad \pi_k^{\text{new}} = \frac{1}{N} \sum_{n=1}^{N} \gamma(z_{nk})$$

5. Compute log-likelihood; if not converged go to 3.



#### The Same Example Again



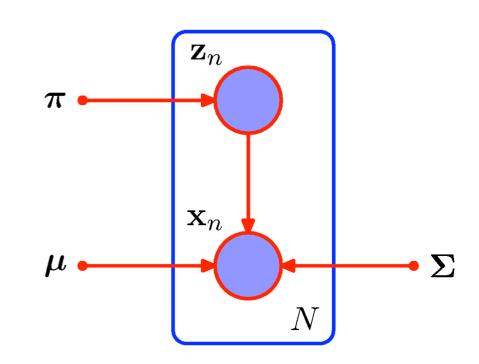
#### Why is it Called "EM"?

Assume for a moment that we observe X and the binary latent variables Z. The likelihood is then:

$$p(X,Z\mid \boldsymbol{\pi},\boldsymbol{\mu},\boldsymbol{\Sigma}) = \prod_{n=1}^{N} p(\mathbf{z}_n\mid \boldsymbol{\pi}) p(\mathbf{x}_n\mid \mathbf{z}_n,\boldsymbol{\mu},\boldsymbol{\Sigma}) \qquad \text{``Complete-data log-likelihood''}$$

where 
$$p(\mathbf{z}_n \mid \boldsymbol{\pi}) = \prod_{k=1}^K \pi_k^{z_{nk}}$$
 and

$$p(\mathbf{x}_n \mid \mathbf{z}_n, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \prod_{k=1}^K \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)^{z_{nk}}$$



which leads to the log-formulation:

$$\log p(X, Z \mid \boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^{N} \sum_{k=1}^{K} z_{nk} (\log \pi_k + \log \mathcal{N}(\mathbf{x}_n \mid \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k))$$